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## Rates of SARS-CoV-2 transmission and vaccination impact the fate of vaccine-resistant strains

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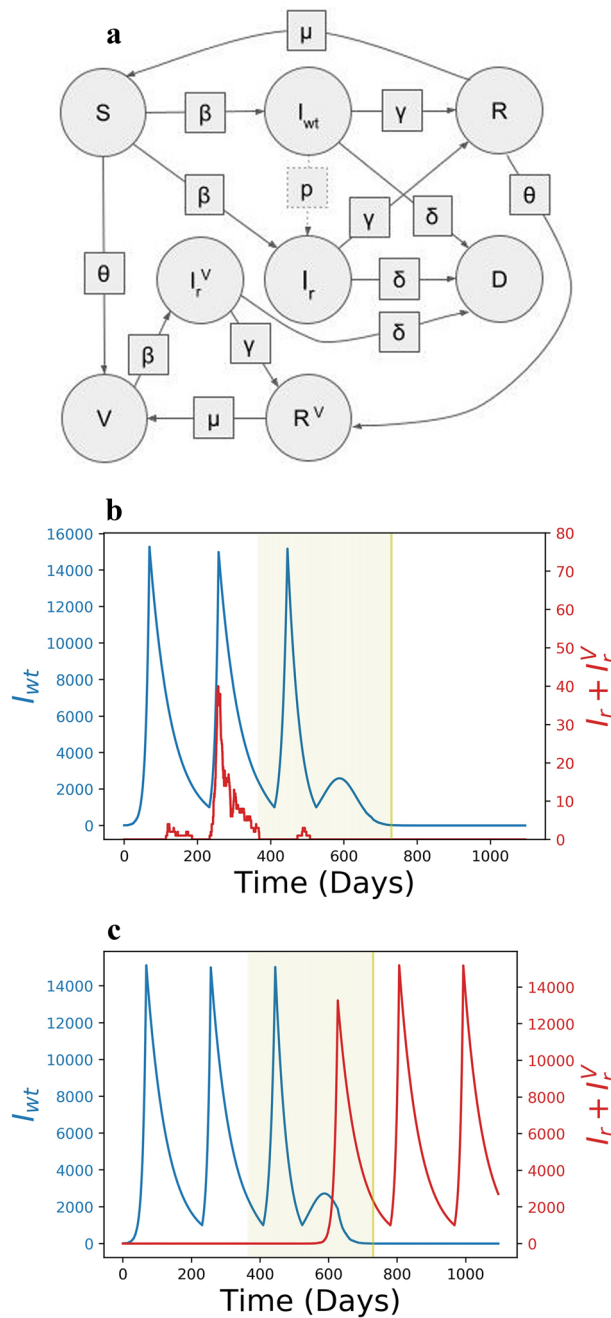
Vaccines are thought to be the best available solution for controlling the ongoing SARS-CoV-2 pandemic. However, the emergence of vaccine-resistant strains may come too rapidly for current vaccine developments to alleviate the health, economic and social consequences of the pandemic. To quantify and characterize the risk of such a scenario, we created a SIR-derived model with initial stochastic dynamics of the vaccine-resistant strain to study the probability of its emergence and establishment. Using parameters realistically resembling SARS-CoV-2 transmission, we model a wave-like pattern of the pandemic and consider the impact of the rate of vaccination and the strength of non-pharmaceutical intervention measures on the probability of emergence of a resistant strain. As expected, we found that a fast rate of vaccination decreases the probability of emergence of a resistant strain. Counterintuitively, when a relaxation of non-pharmaceutical interventions happened at a time when most individuals of the population have already been vaccinated the probability of emergence of a resistant strain was greatly increased. Consequently, we show that a period of transmission reduction close to the end of the vaccination campaign can substantially reduce the probability of resistant strain establishment. Our results suggest that policymakers and individuals should consider maintaining non-pharmaceutical interventions and transmission-reducing behaviours throughout the entire vaccination period.

Vaccines are among the most effective public health measures against infectious disease<sup>1</sup>. Their track record brings hope that SARS-CoV-2 may soon be under control<sup>2</sup> as a consequence of a plethora of vaccine development efforts<sup>3–8</sup>. A potential cause of concern is the low rate of vaccine production and administration<sup>9</sup> coupled with reports of new strains with higher transmission rates<sup>10–12</sup> and even with potential for some degree of vaccine resistance<sup>13–16</sup>. A number of models considered the dynamics of the spread of a vaccine-resistant strain in the population<sup>17–20</sup>. However, to our knowledge, the interplay of the population vaccination rate with the stochastic dynamics of emergence of a resistant strain has been discussed<sup>21</sup>, but not formally modeled. Specifically, a concern is whether a combination of vaccination and transmission rates can create positive selection pressure on the emergence and establishment of resistant strains<sup>22,23</sup>. To address this issue, we implemented a model to simulate the probability of emergence of a resistant strain as a function of vaccination rates and changes in the rate of virus transmission, resembling those caused by non-pharmaceutical interventions and behavioural changes. We then performed a number of simulations based on realistic parameters to study the likelihood and pattern of the emergence of a resistant strain. Finally, we considered possible countermeasures to reduce the probability of the establishment of the resistant strain in the population.

### Results

We implemented a modification of a SIR model<sup>18,24</sup> that included additional states to study the interplay of the rate of vaccination, rate of transmission and the likelihood of emergence of resistant strains (Fig. 1a). In addition to other states, individuals could be vaccinated ( $V$ ), infected by the resistant strain ( $I_r$ ), or simultaneously be vaccinated and infected with the resistant strain ( $I_r^V$ ). The model was run to simulate a population of 10,000,000 individuals over 3 years with the vaccination starting after the first year. In the model, the susceptible individuals ( $S$ ) are infected by the wildtype, or the original, strain at a rate of  $\beta$  and infected individuals recover at a rate of  $\gamma$  or die at a rate of  $\delta$ . At each time step, a fraction  $\theta$  of all non-infected individuals is vaccinated and with some fixed probability  $p$ , an infected individual becomes infected with a resistant strain. Conversely, any individual

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**Figure 1.** The states, transition parameters and dynamics. **(a)** States are shown in circles and transition parameters in squares. The transition parameter,  $\mu$ , is the rate at which individuals lose natural immunity and  $p$ , is the probability that an individual infected with the wildtype strain transmits a resistant strain, so it is not a deterministic parameter. Example dynamics of the number of individuals infected with the wildtype (blue) and resistant strains (red) for  $p = 10^{-6}$ ,  $\theta_0 = 1/365$  and  $F_h = 15,000$ . The period of vaccination is highlighted (green). Under the same parameters the resistant strain may emerge and go extinct, **(b)**, or become established, **(c)**.

infected with the resistant strain can revert back to the wildtype strain population with the same probability,  $p$ . Immunity acquired after infection decayed at a rate of  $\mu$ . Overall, our model included eight character states and six transition parameters between them (Fig. 1a, Table 1).

The rate of transmission in the course of a pandemic is typically cyclical<sup>34–36</sup> due to government interventions<sup>36,37</sup>, behavioural changes<sup>38–40</sup>, and environmental<sup>41,42</sup> and other factors<sup>43,44</sup>. Generally, the number of infected individuals is wave-like, guided by periods of high rate of transmission, followed by periods of a low rate of transmission<sup>31,34,35,45</sup>. We thus varied  $\beta$ , the rate of virus transmission to reflect this cyclical behavior (Fig. 1b,c). A high rate of transmission ( $\beta_h = 0.18$ , equivalent to the effective reproduction number of  $R_h = 2.52$ ) was alternated with a low rate ( $\beta_l = 0.055$  or  $R_l = 0.77$ ), which broadly reflected the observed rates of transmission in various countries affected by the SARS-CoV-19 pandemic with and without lockdown measures,

Parameter	Value	Comments
<b>Fixed parameters</b>		
Population size, $N$	10,000,000 individuals	
Cut-off for stochastic mode, $N^*$	1000	
Cut-off for Gillespie algorithm, $N_c$	100	
Recovery rate, $\gamma$	$0.99 \times 1/14$	Average disease duration is 14 days <sup>25</sup>
Death rate, $\delta$	$0.01 \times 1/14$	Infection-fatality rate is 1% <sup>26–28</sup>
Loss of immunity rate, $\mu$	$1/180$	On average in 180 days <sup>29,30</sup>
Share of non-vaccinated, $h$	0.01	1% of population
Saturation parameter, $k$	0.01	
Transmission rate, $\beta = \{\beta_l, \beta_h\}$	{0.055, 0.18}	$R_0 = \{0.77, 2.52\}$ <sup>31–33</sup>
$\Delta t$	1 day	
<b>Varying parameters</b>		
Bound for initiation of low transmission, $F_h$	2000–20,000 individuals	
Bound for initiation of high transmission, $F_l$	1000 or $F_h/8$ individuals	
Probability of emergence of resistant strain, $p$	$1e-8$ to $1e-5$	Daily for every infected
Vaccination speed, $\theta_0$	0.001 to 0.015	0.1–1.5% daily

**Table 1.** Model parameters.

respectively<sup>31–33</sup>. The low rate of transmission was triggered in the model when the number of individuals infected with any strain reached a high threshold  $F_h = (I_{wt} + I_r + I_r^V)$ . We considered two modes of transition from a low rate of transmission back to a high rate, when the number of infected individuals reached a low threshold  $F_l$  at a constant value of  $F_l = (I_{wt} + I_r + I_r^V) = 1000$ , which was used to generate the main figures, and at a relative value of  $F_l = F_h/8$ , which are available as “Supplementary materials S1”. The explored values of  $F_h$  and  $F_l$  were selected such that the number of infection waves during the first year of the model roughly coincided with the number of waves of SARS-CoV-2 infection observed in most countries during the first year of the pandemic.

SIR-like models frequently consider only deterministic dynamics<sup>18</sup>. However, the emergence of a new strain is an inherently stochastic process under extensive influence of genetic drift<sup>46,47</sup>. Therefore, we incorporated a stochastic stage into our model to allow for genetic drift in the early phases of population dynamics of the resistant strains. The growth rate of the number of individuals infected with the wildtype strain at time  $t$  was determined deterministically by  $(\beta \cdot S/N - \gamma - \delta)I_{wt}$ . By contrast, when the frequency of the resistant strain in the population is low, the number of transmissions of the resistant strain was drawn from a Poisson distribution with a mean of  $(\beta I_r (S + V)/N)dt$ <sup>48</sup>. However, when the frequency of the resistant strain is greater than 1000 individuals (0.01% of the population), making it highly unlikely to disappear by stochastic forces, the dynamics are treated deterministically in the same manner as the wildtype strain.

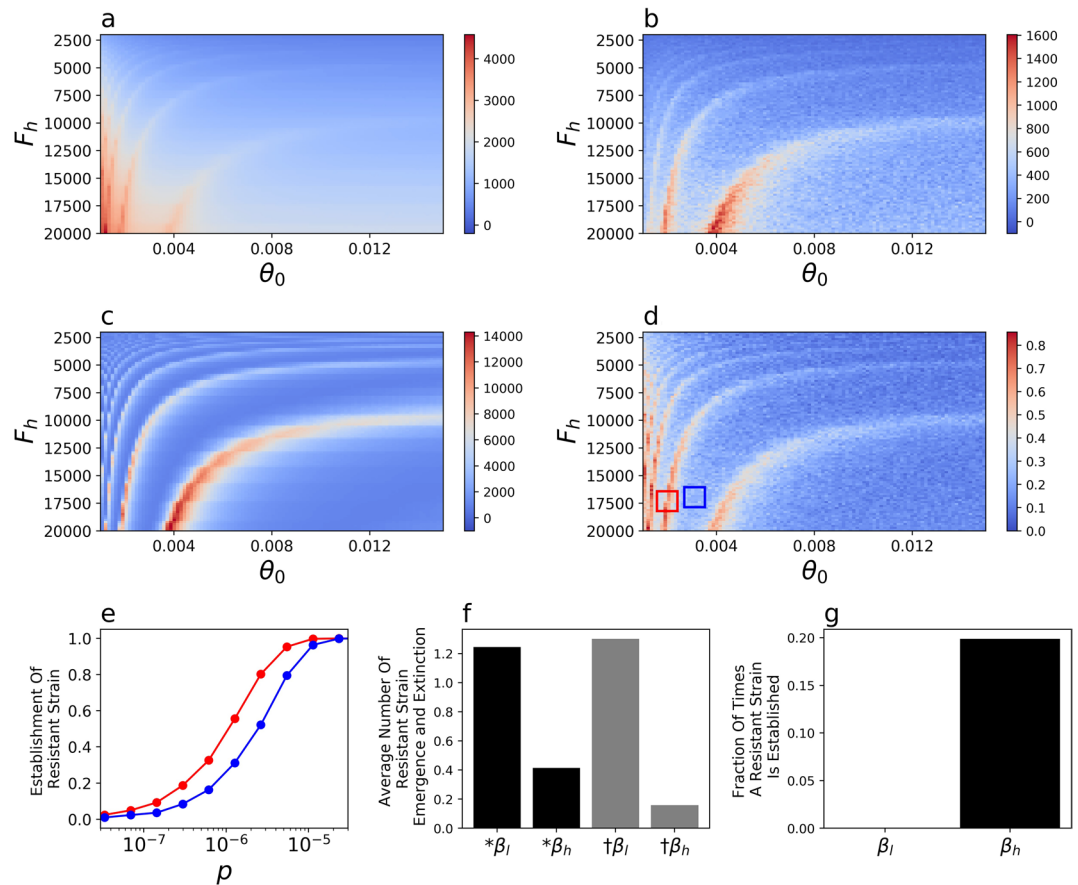
We define three stages of vaccine-resistant strain propagation, including emergence, the appearance of the first individual with the infected strain, establishment, the time point when the number of infected individuals reached 1000, and extinction, when the number of resistant strain infected individuals returned to zero. The impact of three parameters on the resistant strain propagation were explored: the probability of the emergence of the resistant strain ( $p$ ), the speed of vaccination ( $\theta$ ) and the initiation of periods of lower rate of transmission ( $F_h$ ). All other parameters were constants and their values were chosen to be broadly reflecting a realistic set of parameters that approximate the available data for the SARS-CoV-2 pandemic,  $\mu = 1/180$ ,  $\gamma = 0.99 \times 1/14$  and  $\delta = 0.01 \times 1/14$  (see Table 1).

Higher probability of the initial emergence of a resistant strain emerging in a single individual,  $p$ , had a predictably<sup>49</sup> positive effect on the probability of the establishment of the resistant strain (Figs. S1), but depended on the rate of vaccination ( $\theta$ ) and low transmission initiation ( $F_h$ ) in a complex manner (Fig. 2, Figs. S1). The dependency was periodic and the probability of establishment of the resistant strain was different by a factor of two even for some similar values of  $\theta$  and  $F_h$  (Fig. 2d,e).

The behaviour of the emergence, establishment and extinction of the resistant strain in the population bears striking resemblance to the population genetics problem addressing the survival of a beneficial allele in a growing population<sup>50,51</sup>. To understand the stochastic dynamics of the resistant strain in the model, it is therefore instructive to consider the underlying mechanism in population genetics terms. Unless the rate of mutation is zero, or infinitesimally low, new variants will emerge in the population at a rate of  $p \cdot I_{wt}$ . When the rates of transmission of the wildtype and the resistant strain are equal, the probability that the resistant strain will go extinct,  $1 - Q_t$ , can be approximated by

$$1 - Q_t = \exp[-Q_{t+1}(1 + s)R_t^r], \quad (1)$$

where  $s$  is the selection advantage of the resistant strain<sup>46,52</sup> and  $R_t^r = (S + V)\beta/N(\gamma + \delta)$ , or the rate of population growth. Therefore, even when there is no selective advantage of the resistant strain ( $s = 0$ ) over the wildtype strain but the rate of transmission is high ( $R_t^r > 1$ ), the likelihood that a new mutation is lost from the population is small ( $\sim 10\%$  for  $\beta_h = 0.18$ , or  $R_0 = 2.52$ ). By contrast, when the rate of transmission is low ( $R_t^r < 1$ , which is the case during the low transmission periods in the model), the probability of extinction of the resistant strain by

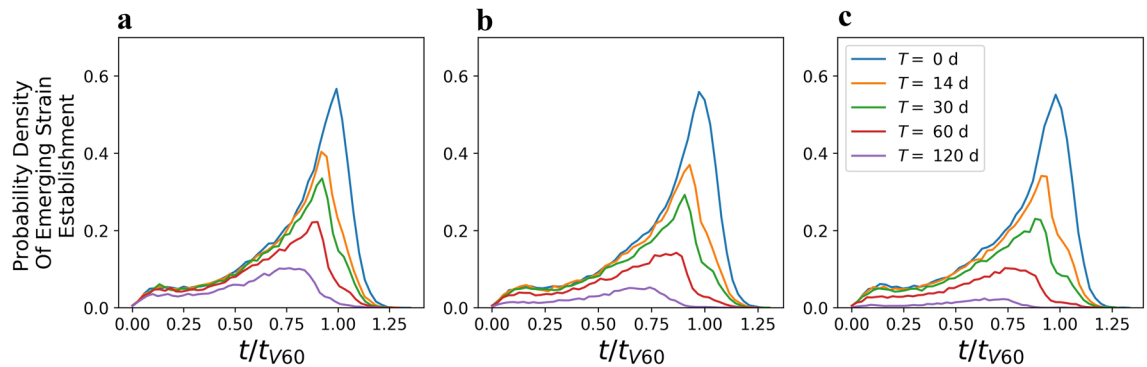


**Figure 2.** Impact of the rate of vaccination,  $\theta$ , and the initiation of low rate of transmission,  $F_h$ , on model dynamics. The cumulative death rate from the (a) wildtype and (b) resistant strains, (c) the number of wildtype-strain infected individuals at  $t_{60}$ , the point in time when 60% of the population is vaccinated and (d) the probability of resistant strain establishment, for  $p = 10^{-6}$ . (e) The probability of emergence of the resistant strain as a function of the probability of emergence,  $p$  shown for the parameter ranges of  $\theta$  and  $F_h$  in the corresponding red and blue boxes from figure (d). (f) The average number of times of  $8 \times 10^6$  simulation runs during which a resistant strain emerges (black) or goes extinct (grey) during periods of low ( $\beta_l$ ) or high ( $\beta_h$ ) transmission for  $p = 10^{-6}$ . (g) A resistant strain was never observed to establish during periods of low transmission ( $\beta_l$ ) for  $p = 10^{-6}$ .

genetic drift is substantial<sup>46,47</sup>. The results of our model are consistent with theory<sup>46,47</sup>, such that the resistant strain emerges during periods of high and low transmission rates, but goes extinct with higher probability during periods of low transmission (Fig. 2f). Furthermore, under the parameters of our model the resistant strain becomes established only when the rate of transmission is high (Fig. 2f).

The complex influence of the speed of vaccination,  $\theta$ , and initiation of low transmission period,  $F_h$ , on the dynamics of establishment of the resistant strain (Fig. 2d) is, therefore, likely driven by the overlap of the vaccination period and the periodicity of the cycles of the number of infected individuals driven by the interaction of  $F_h$  and  $\theta$  (Fig. 1b,c). The coincidence of a high number of vaccinated individuals and a high rate of transmission has two effects on the resistant strain. First, as mentioned previously, because the rate of transmission is high, the emerging resistant strain is not lost through genetic drift (see also<sup>50,53</sup>). Second, a high number of vaccinations creates a selective advantage of the resistant strain over the wildtype strain<sup>23</sup>. The effective reproductive number of the wildtype versus the resistant strains,  $R_t^{wt}/R_t^r$ , is  $(V+S)/S$ , which is the selective advantage  $1+s$  in Eq. (1). Thus, when  $V$  is large the resistant strain has a growth advantage over the wildtype strain, contributing to its establishment in the population towards the end of the vaccination campaign. Taken together, the highest probability for establishment of the resistant strain for a given  $p$  is reached when  $V$ ,  $I_{wt}$  and  $\beta$  (and the corresponding  $R_t^r$ ) are large (Fig. 2c, Eq. (1)).

Indeed, when  $p = 10^{-6}$ , in those cases when the resistant strain becomes established, its initial time of emergence frequently occurs at around the time when 60% of the population is vaccinated (Fig. 3). Therefore, we then tested the influence of a single intervention triggering at a single extraordinary period of low transmission centred around 60% of vaccinated individuals in the population (Fig. 3). We varied the duration of this intervention,  $T$ , ranging from 1 week to 120 days and considered three rates of transmission,  $\beta_l = 0.055$  ( $R_0 = 0.77$ ),



**Figure 3.** Time of initial emergence of a resistant strain that has become established. Probability density that the resistant strain emerges as a function of time since the start of the simulation,  $t$ , rescaled by the time at which 60% of the individuals are vaccinated,  $t_{v60}$ , averaged across simulations with  $\theta$  (0.001 through 0.015),  $F_h$  (2000 through 20,000) and  $p = 10^{-6}$ . Without any extraordinary periods of low transmission (blue line) the peak of the likelihood of emergence of a new strain is at  $t/t_{v60} = 1$ . The likelihood of emergence of a resistant strain can be reduced by an extraordinary period of low transmission centered at  $t/t_{v60} = 1$  with a stronger reduction when such period is longer,  $T$  (colour-coded), or when the rate of transmission is more strongly reduced (a)  $\beta_1 = 0.055$ , (b)  $\beta_1 = 0.03$ , (c)  $\beta_1 = 0.01$ .

0.03 ( $R_0 = 0.42$ ) and 0.01 ( $R_0 = 0.14$ ). Both parameters decrease the probability of establishment of the resistant strain with the length of the intervention having a relatively stronger effect (Fig. 3, Figs. S1, S1, S1, S1, S1, S1, S1).

## Discussion

Our model suggests three specific risk factors that favour the emergence and establishment of a vaccine-resistant strain that are intuitively obvious: high probability of initial emergence of the resistant strain, high number of infected individuals<sup>54</sup> and low rate of vaccination<sup>55</sup>. By contrast, a counterintuitive result of our analysis is that the highest risk of resistant strain establishment occurs when a large fraction of the population has already been vaccinated but the transmission is not controlled. Similar conclusions have been reached in a SIR model of the ongoing pandemic<sup>56</sup> and a model of pathogen escape from host immunity<sup>57</sup>. Furthermore, empirical data consistent with this result has been reported for influenza<sup>58</sup>. Indeed, it seems likely that when a large fraction of the population is vaccinated, especially the high-risk fraction of the population (aged individuals and those with specific underlying conditions) policy makers and individuals will be driven to return to pre-pandemic guidelines<sup>59</sup> and behaviours conducive to a high rate of virus transmission<sup>60,61</sup>. However, the establishment of a resistant strain at that time may lead to serial rounds of resistant strain evolution with vaccine development playing catch up in the evolutionary arms race against novel strains.

Prior to discussion of the implications of our model we reflect on several properties of the assumptions and implementation of our model. In classical SIR-like deterministic models even a single individual infected with a vaccine resistant strain with reproduction number  $R_t > 1$  will lead to automatic establishment of the strain in the population. In an analytical solution, a SIR-like model, even for  $R_t < 1$  for the vaccine resistant strain, the number of infected individuals will tend to 0 but only as time tends to infinity. In actual populations, a single individual infected with a vaccine resistant strain still has a non-negligible chance not to infect anyone causing the variant to go extinct due to random stochastic forces<sup>47</sup>. Therefore, the implementation of stochastic dynamics<sup>62,63</sup> of the vaccine resistant strain at low frequency in our model, considers the impact of random drift on its dynamics, which lies at the heart of extinction of rare strains.

We considered the dynamics of a single vaccine resistant strain, however, there may be different mutations that can lead to vaccine resistance. The emergence of different genotypes causing the same phenotype is analogous to a distinction in population genetics between alleles identical by state and by descent<sup>64</sup>. In our model, the treatment of independent emergence of different mutations as a single entity does not influence the dynamics under the following two assumptions. First, that different mutations lead to exactly the same phenotype, which is vaccine resistance, and, second, that there is no recombination. However, the reported dynamics may be quantitatively different if either of the two assumptions do not hold.

We have not explored the parameter ranges of  $\beta_h$ ,  $\beta_l$ , the high and low rates of transmission, respectively, and  $F_l$  the threshold between low and high rate of transmission. We selected the  $\beta_h$  and  $\beta_l$  to represent the known transmission values at the start of the pandemic<sup>31–33</sup>. However, evolving strains are reported to have a higher rate of transmission<sup>10</sup> leading to higher  $\beta_h$  and, possibly,  $\beta_l$  values than we used. An increase in the rate of transmission is not expected to qualitatively influence the reported dynamics, but would shift the probability density of establishment of a resistant strain (Fig. 3). Indeed, the peak probability at 60% vaccinated individuals roughly corresponds to the point at which for the given  $\beta_h$ ,  $R_p$ , the average number of transmissions for one infected individual, becomes less than 1. Because the reproduction number for the vaccine resistant strain,  $R_p$ , is equal to  $(S + V)\beta/N(\gamma + \delta)$ , the perk of the risk of establishment of the vaccine resistant strain would increase proportional to an increase of  $\beta$ . An increase in either  $F_h$  or  $F_l$  would lead to more individuals becoming infected and a proportionally higher rate of emergence of the vaccine related strain, but would not change the qualitative

behaviour of the model. Furthermore, an increase of  $F_I$  would lead to reversion to a high transmission rate with higher number of infected individuals, leading to shorter periods of low transmission and decreased probability of extinction of the vaccine resistant strains.

The results of our model provide several qualitative implications for the strategy forward in the months of vaccination. In our model, the probability of emergence of a resistant strain in one individual per day was in the range of  $10^{-5}$  to  $10^{-8}$  for a population of  $10^7$  individuals. For the entire human population of  $\sim 10^{10}$  that probability would be  $10^{-8}$  to  $10^{-11}$ , which does not seem improbably large. As of February 2021,  $\sim 10^9$  individuals have been infected by SARS-CoV-2<sup>65</sup> with an average 14 days of sickness per individual<sup>25</sup>, so  $> 10^{10}$  number of total days of infected individuals. Furthermore, highly mutated strains may emerge as a result of long shedding in immunocompromised individuals, a rare but realistic scenario<sup>66–68</sup>. Taken together, the emergence of a partially or fully vaccine-resistant strain and its eventual establishment appears inevitable. However, as vaccination needs to be ahead of the spread of such strains in similar ways to influenza<sup>23</sup>, it is necessary to reduce the probability of establishment by a targeted effort to reduce the virus transmission rate towards the end of the vaccination period before the current vaccines become ineffective. Conversely, lack of non-pharmaceutical interventions at that time can increase the probability of establishment of vaccine-resistant strains. For example, plans to vaccinate individuals with a high risk of a fatal disease outcome followed by a drive to reach herd immunity while in uncontrolled transmission among the rest of the population is likely to greatly increase the probability that a resistant strain is established, annulling the initial vaccination effort. Another potential risk factor may be the reversion of vaccinated individuals to pre-pandemic behaviours that can drive the initial spread of the resistant strain.

One simple specific recommendation is to keep transmission low even when a large fraction of the population has been vaccinated by implementing acute non-pharmaceutical interventions (i.e. strict adherence to social distancing) for a reasonable period of time, to allow emergent lineages of resistant strains to go extinct through stochastic genetic drift. The implementation of non-pharmaceutical measures at a time of high vaccination can also help reduce infectivity when the efficacy of vaccines is not perfect<sup>69</sup>. Additional factors that may make these measures even more effective are: (1) increased and widespread testing, (2) rigorous contact tracing, (3) high rate of viral sequencing of positive cases<sup>58,70</sup> and (4) travel restrictions. Finally, while our model formally considers only one homogenous population, our data also suggest that delays in vaccination in some countries relative to others will make the global emergence of a vaccine-resistant strain more likely. Without global coordination, vaccine resistant strains may be eliminated in some populations but could persist in others. Thus, a truly global vaccination effort may be necessary to reduce the chances of a global spread of a resistant strain.

## Materials and methods

Our extension of the SIR Model features 8 distinct states. Susceptible,  $S$ , and recovered,  $R$ , individuals are vaccinated over time to become vaccinated,  $V$ , or recovered vaccinated,  $R^V$ , respectively. Susceptible individuals can become infected with the wildtype,  $I_{wt}$ , or the resistant virus strain,  $I_r$ . While the vaccinated population is immune to the wildtype, it can be infected by the vaccine-resistant strain, in which case the state is represented by  $I_r^V$ . After a while any infected individual recovers or dies,  $D$ . Finally, we assume that the recovered population retains natural immunity towards both strains, but becomes susceptible again with some small rate,  $\mu$ . In our model, immunity against the wildtype strain gained through vaccination is not lost during the entire model period of 3 years, consistent with current estimates<sup>71,72</sup>.

The total number of individuals,  $N$ , in the population remains constant at 10,000,000, which includes the diseased individuals. We do not introduce new individuals into the population because only a very small number of individuals die during the 3 years that we simulate.

$$S + I_{wt} + I_r + I_r^V + R + R^V + D + V = N \quad (2)$$

For all of the eight states, for convenience we omit the time index, e.g. we write, for example,  $S$  instead of  $S(t)$ . In the limit of large population sizes, the full dynamics without mutations can be described by the following set of differential equations. In these equations,  $\dot{x} = dx/dt$  where  $t$  is time.

$$\dot{S} = \mu R - \theta(t)S - \beta(t)(I_{wt} + I_r + I_r^V)S \quad (3)$$

$$\dot{I}_{wt} = -(\gamma + \delta)I_{wt} + \beta(t)SI_{wt} \quad (4)$$

$$\dot{I}_r = -(\gamma + \delta)I_r + \beta(t)S(I_r + I_r^V) \quad (5)$$

$$\dot{I}_r^V = -(\gamma + \delta)I_r^V + \beta(t)V(I_r + I_r^V) \quad (6)$$

$$\dot{R} = -\mu R - \theta(t)R + \gamma(I_{wt} + I_r) \quad (7)$$

$$\dot{R}^V = -\mu R^V + \theta(t)R + \gamma I_r^V \quad (8)$$

$$\dot{D} = \delta(I_{wt} + I_r + I_r^V) \quad (9)$$

$$\dot{V} = \mu R^V + \theta(t)S - \beta(t)V(I_r + I_r^V) \tag{10}$$

Or, in the matrix format,  $\dot{X} = \hat{P}X$ , with  $X = (S, I_{wt}, I_r, I_r^V, R, R^V, D, V)^T$

$$\hat{P} = \begin{bmatrix} -\theta(t) & -\beta(t)S & -\beta(t)S & -\beta(t)S & \mu & 0 & 0 & 0 \\ \beta(t)I_{wt} & -(\gamma + \delta) & 0 & 0 & 0 & 0 & 0 & 0 \\ \beta(t)I_r & 0 & -(\gamma + \delta) & -\beta(t)S & 0 & 0 & 0 & 0 \\ 0 & 0 & \beta(t)V & -(\gamma + \delta) & 0 & 0 & 0 & \beta(t)I_r^V \\ 0 & \gamma & \gamma & 0 & -(\mu + \theta(t)) & 0 & 0 & 0 \\ 0 & 0 & 0 & \gamma & \theta(t) & -\mu & 0 & 0 \\ 0 & \delta & \delta & \delta & 0 & 0 & 0 & 0 \\ \theta(t) & 0 & -\beta(t)V & -\beta(t)V & 0 & \mu & 0 & 0 \end{bmatrix} \tag{11}$$

The dynamics are influenced by the following constant parameters: the recovery rate,  $\gamma$ , the death rate,  $\delta$ , and the rate at which natural immunity is lost,  $\mu$ . Additionally we introduce a time dependent transmission rate  $\beta(t)$  and a function  $\theta(t)$ , which controls the speed of vaccination.

**Time dependent transmission rate.** Pandemics often proceed in a wave-like pattern<sup>31,34,35,45</sup>, so we introduce a parameter  $\beta(t)$ , which switches between high and low transmission rates. The model begins with a period of a high rate of transmission,  $\beta_h/N$ . A low transmission rate,  $\beta_l/N$ , is initiated when the fraction of individuals infected with any strain,  $I = (I_{wt} + I_r + I_r^V)$ , reaches the value of  $F_h$ . Transition from a period of low to high infection rate occurs at  $F_l=1000$  or  $F_l=F_h/8$ .

**Vaccination.** Vaccination is modelled as almost always a linear function with saturation. The deviation from linearity occurs towards the end of the vaccination period. At that time, there may be fewer individuals that can be vaccinated than the number of individuals vaccinated at any time point, or  $S + R - h < \theta$ . This state can persist for longer than one point in time because in our model infected individuals are vaccinated only once they recover.  $h$  denotes the number of individuals in the population that are never vaccinated. A maximum of  $N - h$  individuals can be vaccinated at the end of the vaccination program. The constant  $k$  controls the saturation of the vaccination speed once the number of susceptible individuals is significantly depleted. The state dependent vaccination speed  $\theta(t)$  is given as:

$$\theta(t) = \left(1 - \frac{h}{S + R + I_{wt} + I_r}\right) \left(\frac{\theta_0}{S + R + k}\right), \tag{12}$$

where  $\theta_0$  can take different values and  $h$  and  $k$  are chosen to be small (see Table 1).

**Integration method.** The deterministic differential equations Eq. (11) were numerically solved using an Euler Forward Integration Scheme, with time step  $\Delta t$ , measured in days.

**Resistant strain.** Each day and for every individual infected with the wildtype strain,  $I_{wt}$ , there is a small probability  $p$ , that a vaccine-resistant strain emerges in that individual. Then this individual switches from state  $I_{wt}$  to state  $I_r$ . Conversely, any individual infected with the resistant strain,  $I_r$ , can revert back to the wildtype strain,  $I_{wt}$ , with the same probability  $p$ . Each time step the number of individuals that transition between being infected with different strains is drawn from a Poisson distribution with mean  $\Delta t p I_{wt}$ , for transition to the resistant strain, or  $\Delta t p I_r$ , for reversion to the wildtype strain. The sum of Poisson distributed random variables is itself a Poisson distributed random variable with the mean corresponding to the sum of the means.

**Stochastic and deterministic regimes.** The population dynamics of a rare variant is an inherently stochastic process<sup>46,47</sup>. We can formally treat the spread of a disease in our model as a stochastic birth–death process. In the following we illustrate this with the number of wildtype infections  $I_{wt}$  as an example. In each infinitesimally small time step  $dt$ , there is a probability  $\beta S I_{wt} dt$ , that the wildtype population  $I_{wt}$  grows by 1,  $I_{wt} \rightarrow I_{wt} + 1$ ,  $I_w$  while the susceptible population is decreased,  $S \rightarrow S - 1$ . Similarly, with probability  $(\gamma + \delta) I_{wt} dt$ ,  $I_{wt}$  is reduced by 1,  $I_{wt} \rightarrow I_{wt} - 1$ , while the number of recovered or dead grows by 1. We carefully model small populations,  $I_{wt} < N^*$ , with  $N^*$  representing a small number of individuals, using a stochastic Tau-Leaping Algorithm<sup>73</sup>. We choose a fixed time step size  $\tau$ , that is equal to the time step of the Euler Integrator  $\Delta t$ . For very small  $I_{wt}$  Tau Leaping Algorithm can produce a negative population<sup>73</sup>. This stems from the fact, that the number of events  $K$  that occur in time  $\tau$  is drawn from a Poisson Distribution, that always assigns a non-zero probability for any  $K > I_{wt}$ . We reduce the chances of such a scenario, by solving the exact SSA Gillespie Algorithm when  $I_{wt}$  is below a critical size  $N^{c48}$ .

For large  $I_{wt}$ , larger than some  $N^*$ , this stochastic process can be approximated with the limiting differential Eq. (4) and an Euler Integration Scheme. Once  $I_{wt} \geq N^*$ , we consider that the resistant strain of the virus is established in the population and we continue modelling it using the deterministic equation<sup>18</sup>.

**Parallel evaluation of deterministic and stochastic variables.** In our model we evaluate deterministic and the stochastic dynamics in parallel. While small populations of infected individuals are treated as stochastic, other variables, such as the number of susceptible individuals,  $S$ , are evaluated within the deterministic regime. While the infection numbers of the wildtype or the emergent strain are in the stochastic regime ( $< N^*$ ),

the corresponding terms that contain the wildtype infections  $I_{wt}$  or the emergent infections  $I_r$  and  $I_r^V$  are removed from the deterministic rate equations. When  $I_{wt}$  or  $I_r + I_r^V$  grow above the threshold value  $N^*$ , the corresponding population of infected individuals is treated as deterministic.

**Sources of errors.** Finally, we discuss some sources of errors in our simulation: (1) depending on the time step  $\Delta t$  the Euler Integration Scheme is not exact. In most of our simulations, we choose a time step of 1 day,  $\Delta t = 1d$ . (2) Using the deterministic rate equations for the infection numbers in Eq. (11) is an approximation to the exact stochastic dynamics given by a birth–death model. The quality of this approximation is given by the threshold value  $N^*$ , which was 1000 in our model. (3) When  $I_{wt}$  or  $I_r + I_r^V$  trespasses the threshold  $N^*$  from above, the populations of infected individuals changes from being treated as a real number (the mean field average) to being treated as a natural number. We truncate the mean field average with a floor function and treat the remainder as part of the recovered population. (4) The Tau Leaping algorithm is an approximation to the exact SSA Gillespie Algorithm, that allows faster evaluation with a constant step size  $\tau$ . Increasing the threshold value  $N_c$  increases the accuracy of the model (Fig. S1). (5) As discussed above, on rare occasions a population of infected individuals drops below 0 in one leap. If this happens we redraw from the same Poisson distribution. (6) Finally, while the time step of the deterministic model  $\Delta t$  and  $\tau$  are chosen to be equal, for population sizes above  $N_c$ , the SSA algorithm acts on exponentially distributed waiting times  $\tau_{SSA}$  between reactions<sup>48</sup>. This introduces errors, if  $\tau_{SSA} \gg \Delta t$ , because the interaction rates may change, while we wait for the next reaction to occur.

In order to determine a range of acceptable values of  $N_c$ , we ran our simulation for a period of  $T = 200$  days, initially loading the system with  $I_{wt} = 200$  wildtype carriers. For multiple values of  $N_c$  and no mutations, we compared the results of disease survival with the analytical solution derived for the birth death process,

$$P(I_{wt}(T) = 0) = 1 - \left( \frac{(\delta + \gamma)e^{(\delta + \gamma - \beta)T} - \gamma - \delta}{(\delta + \gamma)e^{(\delta + \gamma - \beta)T} - \beta} \right)^{I_{wt}(0)}. \quad (13)$$

We pick  $N_c = 100$  individuals, which gives us a reasonably small error.

**Assumptions and choice of parameters.** The model is run for a total time of 3 years, with vaccination starting 1 year into the model. We assume that the wildtype and emergent strains have the same infectivity ( $\beta$  is the same for both strains). We assume that infection by any one strain provides immunity to both, reflecting that many vaccines carry only the Spike protein of the SARS-CoV-2 virus and it may be easier to escape immunity provided by the vaccine than the immunity provided by infection. We also assume that the immune response provided by the vaccine is more permanent and that immunity provided by infection, is lost at rate  $\mu$ , on average after 0.5 years<sup>2,71,72,74</sup> after recovery. Both of these assumptions influence the model when the number of infected individuals becomes large, which is unlikely for realistic average rates of transmission across the simulated time.

We assume that susceptible and recovered individuals have an equal chance to be vaccinated,  $\theta_0$ . We also assume that the infection-recovery rate,  $\gamma$ , and infection-fatality rate,  $\delta$  are the same for the wildtype and mutated strains.

We regulate the rates of transmission exogenously in the model, with the rate of transmission ( $\beta$ ) switching between a high rate,  $\beta_h$  and a low rate,  $\beta_l$ , when the total number of individuals infected with either strain reaches a threshold parameter. These threshold parameters,  $F_h$  and  $F_l$ , simultaneously reflect the impact of all non-pharmaceutical interventions and behavioural changes of individuals. The ranges of these parameters were chosen to broadly reflect realistic parameters of a pandemic, including rates of infection and several waves of high infection in the first couple of years of the model.

We operate under the assumption that vaccine efficacy not only impacts disease manifestation but also blocks transmission at the same rate, which is a reasonable assumption based on previous vaccine performance but has not yet been demonstrated.

In Table 1 we present the choice of parameters for the model, the ones that were constant and those that were varied, including their boundaries.

## Code availability

<https://github.com/Simon-Re/mut-vacc>.

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## Author contributions

E.T.D. and F.A.K. formulated the question. S.A.R. and Y.A.K. formulated the model. S.A.R. programmed and ran the model. All authors jointly analyzed the results and wrote the manuscript.

## Competing interests

The authors declare no competing interests.

## Additional information

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